Prevalence of Kidney Stones in the United States

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1. Introduction

Urinary stone disease is common and poses a significant health care burden in a working-age population. In 1994, data from the National Health and Nutrition Examination Survey (NHANES) estimated the prevalence of stone disease at 6.3% among men and 4.1% among women [1]. More recent analyses of claims data suggest rising utilization of health care resources for treating patients with stone disease [2,3]. However, there has been no nationally...
representative assessment of the prevalence of kidney stones in the United States since 1994. Thus, it remains unclear whether the claims data reflect changing patterns of care or changes in the epidemiology of kidney stones.

Recent investigations suggest that dietary and lifestyle factors play an important role in the risk of stone disease. Prospective data from the Health Professionals Follow-up Study and the Nurses’ Health Study I and II note associations between incident stone disease and weight gain, body mass index (BMI), and diabetes mellitus \cite{4,5}. These findings, combined with an obesity epidemic \cite{6}, support the hypothesis that the epidemiology of stone disease in the United States is changing.

Given this context, we sought to describe the current epidemiology of stone disease using a population-based cohort. We used data from the 2007–2010 NHANES sample to estimate the prevalence of stone disease in the United States. In the context of dramatically increasing obesity, we hypothesized that the prevalence of stone disease would be higher than 13 yr previously, when these data were last queried by NHANES III.

2. Methods

2.1. Data source

We used data from the 2007–2010 NHANES. NHANES is a stratified, multistage probability survey of the noninstitutionalized US population. The institutional review board of RAND Corporation determined that the study design was exempted from the requirement for review.

2.2. Study population

The study population consisted of all participants responding to the questions regarding a history of kidney stones contained in the household survey component. These data were used to determine the population prevalence of stone disease. The participants in the medical examination component were used to identify risk factors associated with a history of kidney stones.

2.3. Outcomes

The primary outcome for the analysis was the response to the question, “Have you ever had kidney stones?” Since some stones may be incidental findings on abdominal imaging, our secondary outcome was based on participant response to the question, “How many times have you passed a kidney stone?” We considered any subject who reported passing at least one stone as having a history of symptomatic stone disease.

2.4. Covariates

Subject age and gender were used as reported in the public-use NHANES data. We used the racial and ethnic group variable reported by NHANES based on survey responses to questions about racial group and ethnicity. We combined responses from the NHANES categories of Mexican-American and Other Hispanic into a single Hispanic category according to analytic guidelines \cite{7}. Otherwise, we used the racial/ethnic groups as reported by NHANES. We included household income as a measure of socioeconomic status, categorized as $0–19 999, $20 000–34 999, $35 000–74 999, and ≥$75 000. A personal history of diabetes or gout was identified from the medical examination component, as was BMI, which was categorized as normal weight (18.0–24.9 kg/m\(^2\)), overweight (25.0–29.9 kg/m\(^2\)), and obese (≥30.0 kg/m\(^2\)).

2.5. Statistical analysis

We calculated the percent prevalence of a self-reported history of stone disease, using weights and design variables as reported by NHANES to account for the complex survey structure and produce estimates representative of the total civilian, noninstitutionalized US population. We also adjusted prevalence estimates using data from the 1990 census to account for population changes since previous NHANES items regarding kidney stones were queried. Overall, changes in estimates because of population adjustment were small; therefore, we report unadjusted figures except where otherwise noted. A logistic regression model was constructed to identify factors associated with prevalent stone disease and account for the complex survey structure. We used SAS v.9.2 (Cary, NC, USA) for all analyses. Results were considered statistically significant with two-sided \(p = 0.05\).

3. Results

A total of 12 110 participants responded to the question regarding a history of stone disease. The weighted overall prevalence of stone disease was 8.8% (95% confidence interval [CI], 8.1–9.5). After population adjustment, the prevalence of stone disease was 8.4% (95% CI, 7.7–9.0). Men were more likely to report a history of stone disease than women (Table 1). Overall, estimates of the prevalence of kidney stones decreased only slightly after population adjustment (Table 2); therefore, we report unadjusted prevalence estimates unless otherwise specified. In men

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>History of kidney stones, % (95% CI)</th>
<th>History of passing at least one kidney stone, % (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Male</td>
<td>Female</td>
</tr>
<tr>
<td>Age group, yr</td>
<td></td>
<td></td>
</tr>
<tr>
<td>20–29</td>
<td>3.4  (2.1–4.7)</td>
<td>3.4  (2.2–4.7)</td>
</tr>
<tr>
<td>30–39</td>
<td>6.9  (5.0–8.8)</td>
<td>5.9  (4.5–7.2)</td>
</tr>
<tr>
<td>40–49</td>
<td>9.8  (7.3–12.3)</td>
<td>7.6  (5.6–9.5)</td>
</tr>
<tr>
<td>50–59</td>
<td>13.1 (10.3–15.9)</td>
<td>8.1  (5.3–10.3)</td>
</tr>
<tr>
<td>60–69</td>
<td>19.1 (15.9–22.4)</td>
<td>9.4  (6.6–12.2)</td>
</tr>
<tr>
<td>70+</td>
<td>18.8 (16.5–21.0)</td>
<td>9.4  (7.5–11.3)</td>
</tr>
<tr>
<td>All ages</td>
<td>10.6 (9.4–11.9)</td>
<td>7.1  (6.4–7.8)</td>
</tr>
</tbody>
</table>

CI = confidence interval.
and women, the weighted prevalence increased with age; these effects were observed among all racial/ethnic groups (Table 3). Similar results were seen for the secondary outcome of symptomatic stone disease, although the point estimates were slightly lower (Table 1).

Variation in the prevalence of stone disease existed among different racial/ethnic groups (Table 3). The prevalence of stone disease was highest among non-Hispanic white individuals, at 10.3% (95% CI, 9.4–11.1). Other ethnic/racial groups had a lower prevalence of stone disease. Among Hispanic individuals, the prevalence of kidney stones was 6.4% (95% CI, 5.6–7.3), and among black, non-Hispanic individuals, the prevalence was 4.3% (95% CI, 3.4–5.3). This variation persisted after population adjustment.

BMI was associated with a history of stone disease (Fig. 1). The prevalence of kidney stones was higher among obese individuals (11.2% [95% CI, 10.0–12.3]) and overweight individuals (9.2% [95% CI, 7.9–10.5]) than among individuals of normal weight (6.1% [95% CI, 4.8–7.4]) (p < 0.001). For males and females, respondents at normal weight had the lowest prevalence of stone disease (males: 7.8% [95% CI, 5.4–10.1]; females: 4.4% [95% CI, 3.8–6.1]). Among obese males the prevalence of stone disease was 13.0% (95% CI, 11.0–15.1), and among obese females the prevalence was 9.6% (95% CI, 8.3–10.8).

### Table 2 – Unadjusted and population-adjusted percent prevalence of stone disease by population characteristic

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>History of kidney stones, males</th>
<th>History of kidney stones, females</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Unadjusted, % (95% CI)</td>
<td>Adjusted, % (95% CI)</td>
</tr>
<tr>
<td>All groups</td>
<td>10.6 (9.4–11.9)</td>
<td>10.3 (9.2–11.3)</td>
</tr>
<tr>
<td>Non-Hispanic, white</td>
<td>12.8 (11.3–14.3)</td>
<td>11.8 (10.4–13.2)</td>
</tr>
<tr>
<td>Hispanic</td>
<td>7.1 (5.7–8.4)</td>
<td>8.8 (7.4–10.2)</td>
</tr>
<tr>
<td>Non-Hispanic, black</td>
<td>4.5 (3.4–5.6)</td>
<td>4.8 (3.7–5.9)</td>
</tr>
<tr>
<td>Other race/multiracial</td>
<td>5.6 (2.5–8.8)</td>
<td>5.3 (2.2–8.5)</td>
</tr>
<tr>
<td></td>
<td>7.1 (6.4–7.8)</td>
<td>6.7 (6.1–7.4)</td>
</tr>
<tr>
<td></td>
<td>7.9 (7.0–8.8)</td>
<td>7.5 (6.7–8.4)</td>
</tr>
<tr>
<td></td>
<td>5.7 (4.6–6.9)</td>
<td>6.1 (4.9–7.3)</td>
</tr>
<tr>
<td></td>
<td>4.2 (2.7–5.7)</td>
<td>4.2 (2.8–5.6)</td>
</tr>
<tr>
<td></td>
<td>6.1 (2.7–9.6)</td>
<td>5.6 (2.4–8.8)</td>
</tr>
</tbody>
</table>

CI = confidence interval.

### Table 3 – Weighted (unadjusted) percent prevalence of stone disease by age and race

<table>
<thead>
<tr>
<th>Age group, yr</th>
<th>History of kidney stones, % (95% CI)</th>
<th>History of passing at least one kidney stone, % (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>White (non-Hispanic)</td>
<td>Black (non-Hispanic)</td>
</tr>
<tr>
<td>20–39</td>
<td>6.1 (4.5–7.7)</td>
<td>2.2 (1.3–3.0)</td>
</tr>
<tr>
<td>40–59</td>
<td>10.6 (8.9–12.3)</td>
<td>4.8 (3.0–6.6)</td>
</tr>
<tr>
<td>60–74</td>
<td>14.4 (12.4–16.3)</td>
<td>7.5 (5.5–9.5)</td>
</tr>
<tr>
<td>≥75</td>
<td>15.2 (12.9–17.6)</td>
<td>10.5 (4.2–16.8)</td>
</tr>
<tr>
<td>All ages</td>
<td>10.3 (9.4–11.1)</td>
<td>4.3 (3.4–5.3)</td>
</tr>
</tbody>
</table>

CI = confidence interval.
Several respondent characteristics were associated with a history of kidney stones in the multivariable model (Table 4). The odds of kidney stones in females was approximately two-thirds that in men (odds ratio [OR]: 0.63 [95% CI, 0.52–0.75]; p < 0.001). Individuals reporting their race/ethnicity as Hispanic or black, non-Hispanic were also less likely to report a history of stone disease (Hispanic OR: 0.60 [95% CI, 0.49–0.73], p < 0.001; black, non-Hispanic OR: 0.37 [95% CI, 0.28–0.49], p < 0.001).

Socioeconomic status was also associated with a history of kidney stones. Those individuals with an annual household income $<19,999 were more likely than individuals earning $>75,000 annually to report a history of kidney stones (OR: 1.57 [95% CI, 1.17–2.09], p = 0.002). Individuals with an annual household income between $20,000 and $34,999 showed a similar increase in risk (OR: 1.65 [95% CI, 1.27–2.15], p < 0.001) compared with the high-income group.

Conditions associated the metabolic syndrome were also associated with a reported history of stone disease (Table 4). The odds of kidney stones were 1.55 times higher among obese participants than among individuals of normal weight (OR: 1.55 [95% CI, 1.25–1.94], p < 0.001). Respondents with a history of diabetes were also more likely to report kidney stones (OR: 1.59 [95% CI, 1.22–2.07], p < 0.001). A statistically significant relationship among dietary intake variables and history of kidney stones was observed only for high levels of sodium intake (data not shown). For these reasons, we did not include dietary intake variables in the final model (Table 4).

4. Discussion

In the first NHANES assessment of kidney stone disease since 1994, we found that 1 in 11 persons in the United States reports a history of stone disease. Among the highest-risk groups, the prevalence of stone disease approached one in five persons. These findings represent a marked increase in the prevalence of stone disease over the intervening 15 years: In 1994, the reported prevalence of stone disease in the NHANES sample was only 5.2% (ie, 1 in 20 persons) [1]. This dramatic increase in the prevalence of stone disease persisted even after adjusting for the changes in the US population since NHANES III.

The changing epidemiology of kidney stone disease was not limited to one specific demographic cohort; increasing prevalence of stones was noted in all age, gender, and racial/ethnic groups, even on a population-adjusted basis. Among men, the overall prevalence of stone disease increased from 6.3% in 1988–1994 to a population-adjusted 10.3% in 2007–2010, representing a 63% relative increase in the prevalence of stone disease [1]. The change in prevalence of kidney stones was particularly dramatic in black, non-Hispanics (a 1.7% prevalence in 1988–1994 [1] compared with a population-adjusted 4.5% prevalence in 2007–2010, representing a relative increase of >150%). The relative prevalence increase among Hispanic individuals was also dramatic.

Recent claims-based analyses of health care utilization, particularly in the inpatient setting, have suggested a narrowing of the risk difference for stone disease between men and women [2,3]. Our data show that the overall prevalence of stone disease is 10.6% in males and 7.1% in females. At least two potential explanations exist for these apparently contradictory findings. First, these different findings could represent gender-based practice variation; that is, for reasons that are not illuminated by these data, men may be less likely to be treated as inpatients for stone disease. Alternately, our data are potentially consistent with the hypothesis that in the past, men were much more likely to form stones (represented by the large difference in prevalence between males and females aged ≥50 yr), but risks are currently more equal, as represented by the very similar prevalence of stone disease among males and females aged 20–49 yr. Because of the cross-sectional nature of the NHANES dataset, it is difficult to conclusively test this hypothesis. The NHANES dataset is specifically designed to estimate the prevalence of conditions in the US population and should be considered more reliable than claims-based utilization data for addressing questions of disease prevalence.

Prior analyses of NHANES data did not explore associations between markers of the metabolic syndrome (eg, overweight/obesity, diabetes) and the prevalence of stone disease. Since the last NHANES analysis, however, prospective data from three large cohorts clearly demonstrate an increasing risk of incident stone disease with obesity, weight gain, and other measures related to the metabolic syndrome [4,5]. Physiologically, obesity has been linked to increased renal excretion of calcium and uric acid, as well as increased urine acidity, all of which increase the risk of stone formation [8–13]. For these reasons, we investigated potential associations between obesity and a history of kidney stones. In multivariable models, we observed statistically significant relationships between a history of kidney stones and gout, diabetes, and obesity.
Presuming obesity as a marker for the metabolic syndrome, which is linked epidemiologically and physiologically to risk of kidney stones, the epidemic of obesity in the United States is a likely explanation for the dramatic rise in the prevalence of stone disease. In our analysis, obesity was an equalizer between sexes; the prevalence of stone disease among obese women was higher than in normal-weight men. In addition, the rates of overweight/obesity are higher among black, non-Hispanic and Hispanic individuals than among white, non-Hispanic individuals [14]. These physiologic and epidemiologic data provide a plausible explanation for the observed changes in prevalence of stone disease in our sample compared with results from NHANES III [1].

Other factors may also contribute to the observed changes in stone disease prevalence. Mean average temperature is believed to be a major contributor to variation in geographic risk for stone disease [15,16]. Data from military deployments suggest that the onset of temperature-related changes in stone risk may occur within 90 d [17–20]. For these reasons, Brikowski and colleagues [21] speculate that temperature-related changes have contributed to, and may continue to exacerbate, the rising burden of stone disease.

Our findings should be considered in the context of important limitations. The rapid increase in radiologic imaging in the United States could translate into an increased diagnosis of asymptomatic kidney stones. To examine this possibility, we analyzed self-report of kidney stone passage as a secondary outcome. While the prevalence of stone passage was slightly lower than self-report of a history of kidney stones, these estimates still represent a marked increase in stone disease. The prevalence of stone disease increases with age; therefore, the aging of the US population would be expected to contribute to an overall increase in the prevalence of stone disease absent any other influences. However, population adjustment of our estimates suggested that changes in prevalence because of the aging of the population only partially explained the observed changes in prevalence.

Finally, our results should be considered in the context of the cross-sectional design of the NHANES instrument. While we explored nutritional intake (including moisture) in our multivariable models, we did not find important associations with a history of stone disease. This finding is consistent with the findings from the NHANES III analysis, in which significant associations between dietary factors and stone history were not observed [1]. Since the temporal relationship of participant diet at the time of NHANES and a past history of stone disease is unclear, it is not surprising that prospectively identified associations between diet and incident stone disease [4,5] were not observed in this cross-sectional design.

These limitations notwithstanding, our findings have important implications for treating patients with kidney stones in the United States. Continued increases in prevalence will presumably sustain recent increases in health care expenditures for treating patients with stone disease [2,22]. Currently, patient treatment is characterized primarily by surgical intervention for symptomatic stone events rather than management of the chronic metabolic condition that leads to recurrence in >50% of patients within 5 yr in a primarily working-age population [23]. Evidence from randomized controlled trials suggests that lifestyle and pharmacologic interventions are effective for secondary prevention of kidney stones [24–27]. The rapidly changing prevalence of kidney stones suggests that further efforts toward secondary prevention may be indicated.

5. Conclusions

The prevalence of kidney stones in the United States is markedly higher than when last measured in 1988–1994, and this increase is likely related to dietary and lifestyle factors. Given the temporary disability imposed by a symptomatic stone event, these findings have important implications for a disease that burdens a primarily working-age population. These findings suggest that the direct and indirect costs of kidney stones will continue to rise in the United States, and efforts should be directed toward ameliorating the burden of urinary stone disease.

Author contributions: Charles D. Scales Jr. had full access to all the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis.

Study concept and design: Scales, Saigal.
Acquisition of data: Smith, Hanley, Saigal.
Analysis and interpretation of data: Scales, Smith, Hanley, Saigal.
Drafting of the manuscript: Scales.
Critical revision of the manuscript for important intellectual content: Smith, Hanley, Saigal.
Statistical analysis: Scales, Smith, Hanley.
Obtaining funding: Saigal.
Administrative, technical, or material support: Smith, Hanley, Saigal.
Supervision: Saigal.
Other (specify): None.

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References


